

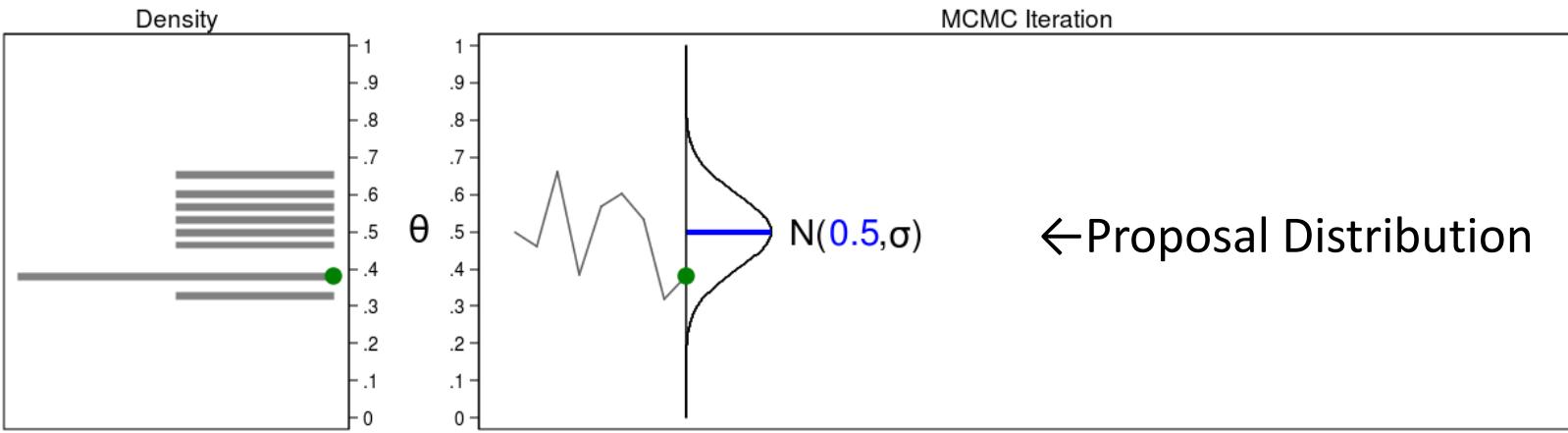
Markov Chain Monte Carlo

Often the posterior distribution does not have a simple form. We can use Markov Chain Monte Carlo (MCMC) with the Metropolis-Hastings algorithm to generate a sample from the posterior distribution.

MCMC and Metropolis-Hastings

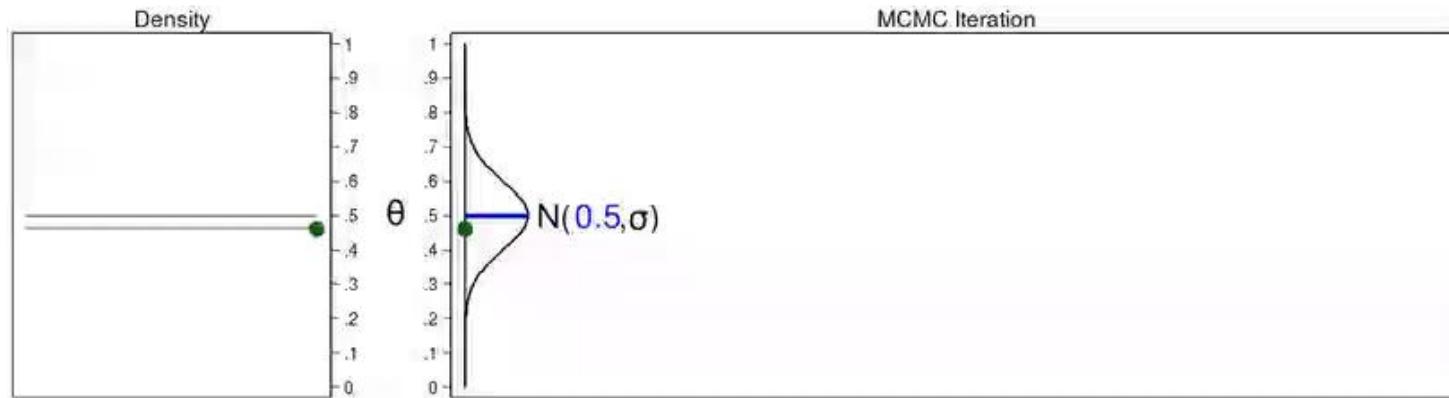
1. Monte Carlo
2. Markov Chains
3. Metropolis-Hastings

Monte Carlo



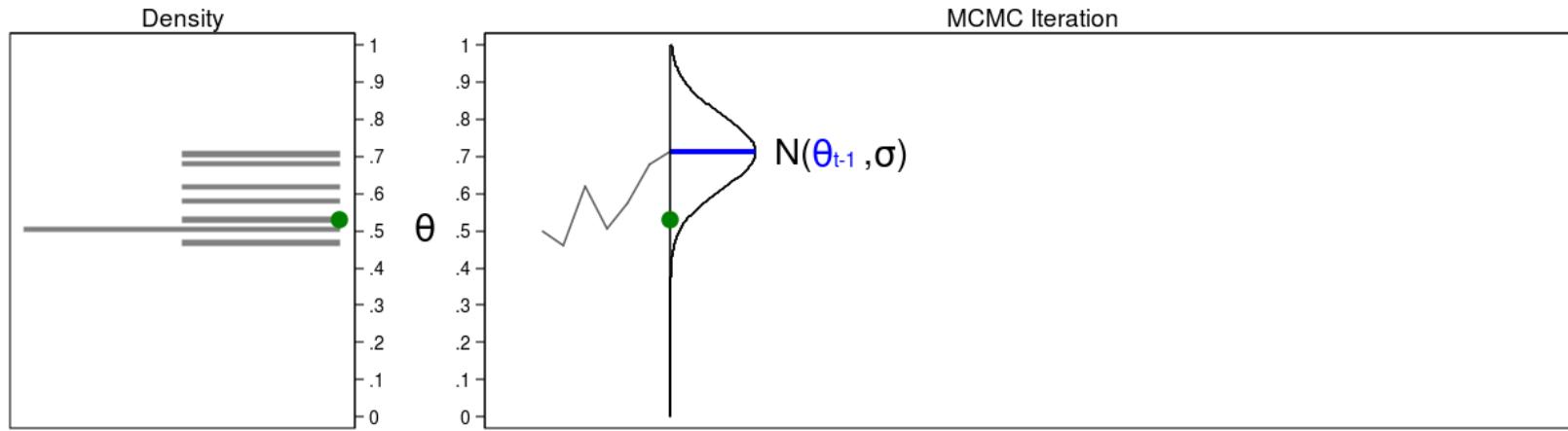
Draw $\theta_t \sim \text{Normal}(0.5, \sigma) = 0.381$

Monte Carlo



Draw $\theta_t \sim \text{Normal}(0.5, \sigma) = 0.460$

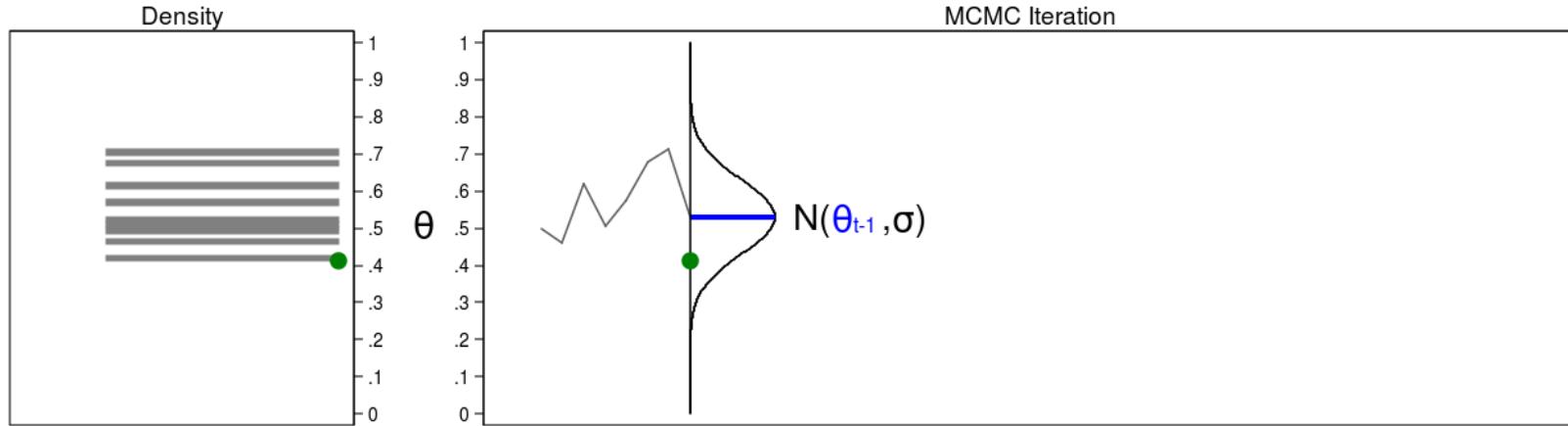
Markov Chain Monte Carlo



Draw $\theta_t \sim \text{Normal}(\theta_{t-1}, \sigma)$

$$\text{Normal}(0.712, \sigma) = 0.530$$

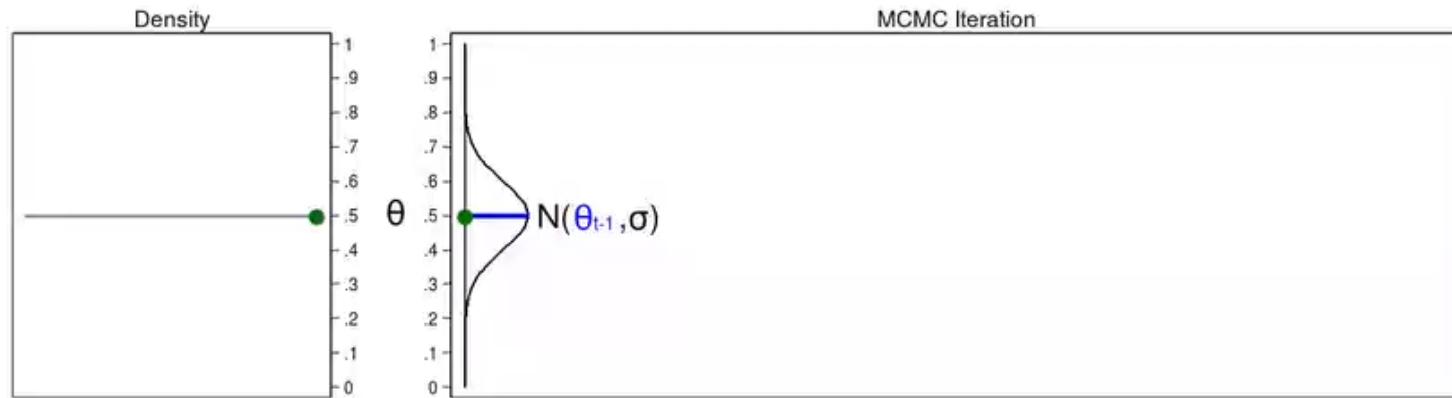
Markov Chain Monte Carlo



Draw $\theta_t \sim \text{Normal}(\theta_{t-1}, \sigma)$

$\text{Normal}(0.530, \sigma) = 0.411$

Markov Chain Monte Carlo



Draw $\theta_t \sim \text{Normal}(\theta_{t-1}, \sigma)$

Normal(0.500, σ) = 0.497

MCMC with Metropolis-Hastings

$$r(\theta_{new}, \theta_{t-1}) = \frac{\text{Posterior probability of } \theta_{new}}{\text{Posterior probability of } \theta_{t-1}}$$
$$= \frac{\text{Beta}(1,1, \theta_{new}) \times \text{Binomial}(10,4, \theta_{new})}{\text{Beta}(1,1, \theta_{t-1}) \times \text{Binomial}(10,4, \theta_{t-1})}$$

MCMC with Metropolis-Hastings

$$\begin{aligned} \text{acceptance probability} &= \alpha(\theta_{new}, \theta_{t-1}) \\ &= \min[r(\theta_{new}, \theta_{t-1}), 1] \end{aligned}$$

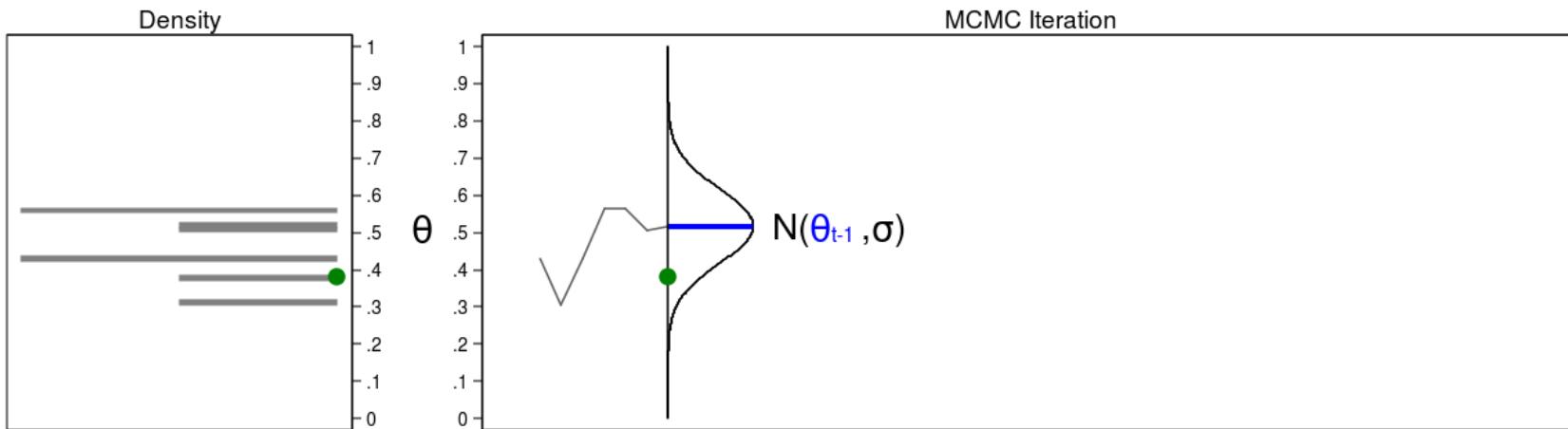
MCMC with Metropolis-Hastings

Draw $u \sim Uniform(0,1)$

If $u < \alpha(\theta_{new}, \theta_{t-1})$ Then $\theta_t = \theta_{new}$

Otherwise $\theta_t = \theta_{t-1}$

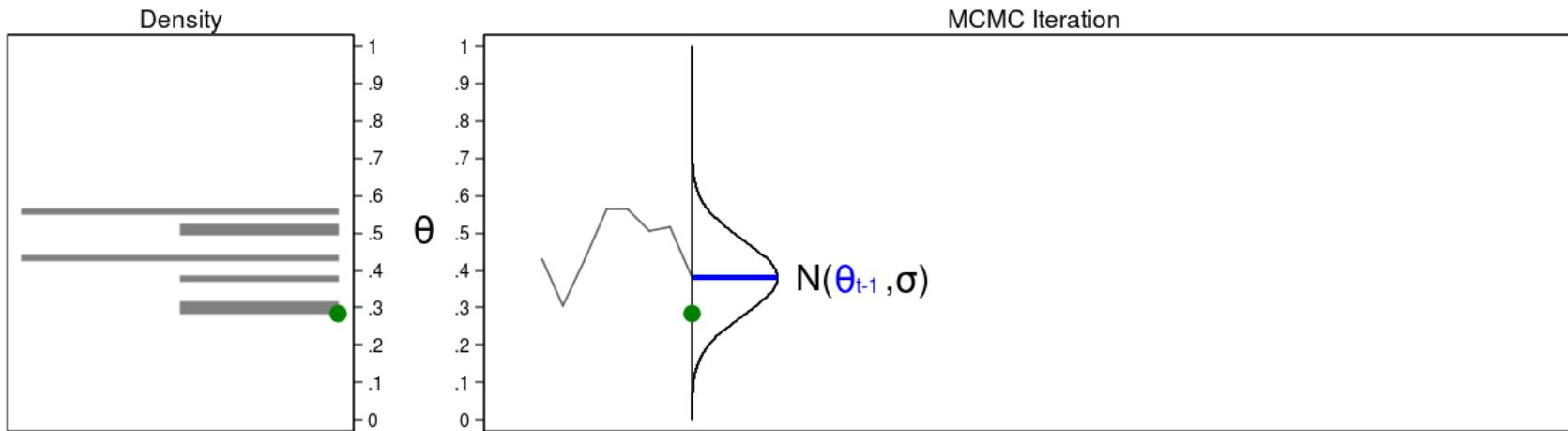
MCMC with Metropolis-Hastings



$$\text{Step 1: } r(\theta_{\text{new}}, \theta_{t-1}) = \frac{\text{Posterior}(\theta_{\text{new}})}{\text{Posterior}(\theta_{t-1})} = \frac{\text{Beta}(1,1, 0.380) \times \text{Binomial}(10,4, 0.380)}{\text{Beta}(1,1, 0.517) \times \text{Binomial}(10,4, 0.517)} = 1.307$$

$$\text{Step 2: Acceptance probability } \alpha(\theta_{\text{new}}, \theta_{t-1}) = \min\{r(\theta_{\text{new}}, \theta_{t-1}), 1\} = \min\{1.000, 1\} = 1.000$$

MCMC with Metropolis-Hastings



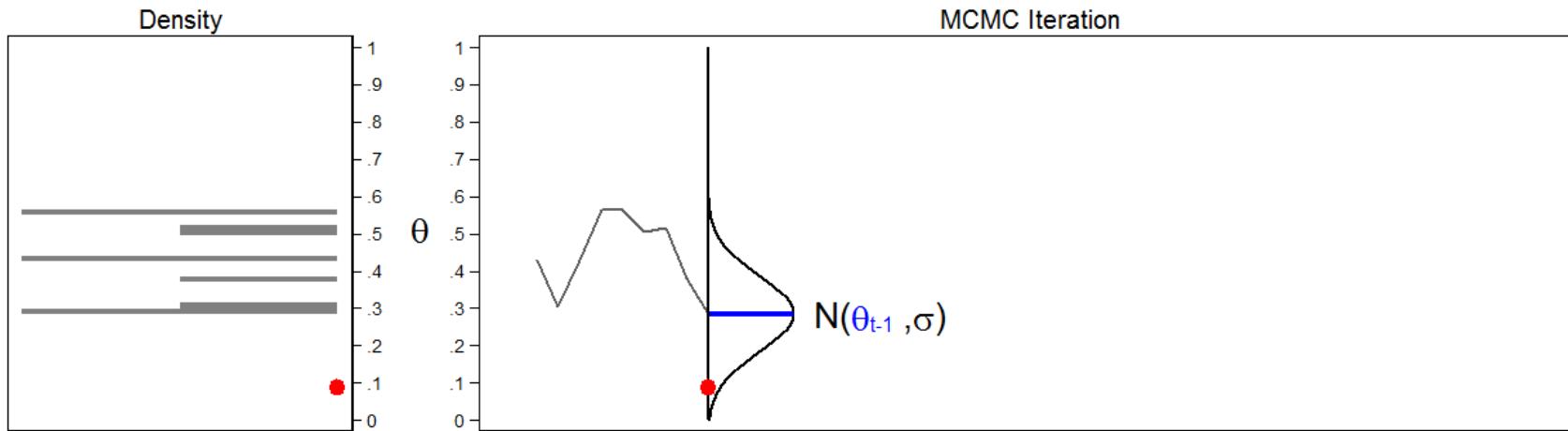
$$\text{Step 1: } r(\theta_{\text{new}}, \theta_{t-1}) = \frac{\text{Posterior}(\theta_{\text{new}})}{\text{Posterior}(\theta_{t-1})} = \frac{\text{Beta}(1,1, 0.286) \times \text{Binomial}(10,4, 0.286)}{\text{Beta}(1,1, 0.380) \times \text{Binomial}(10,4, 0.380)} = 0.747$$

$$\text{Step 2: Acceptance probability } \alpha(\theta_{\text{new}}, \theta_{t-1}) = \min\{r(\theta_{\text{new}}, \theta_{t-1}), 1\} = \min\{0.747, 1\} = 0.747$$

Step 3: Draw $u \sim \text{Uniform}(0,1) = 0.094$

Step 4: If $u < \alpha(\theta_{\text{new}}, \theta_{t-1}) \rightarrow$ If $0.094 < 0.747$ Then $\theta_t = \theta_{\text{new}} = 0.286$
Otherwise $\theta_t = \theta_{t-1} = 0.380$

MCMC with Metropolis-Hastings



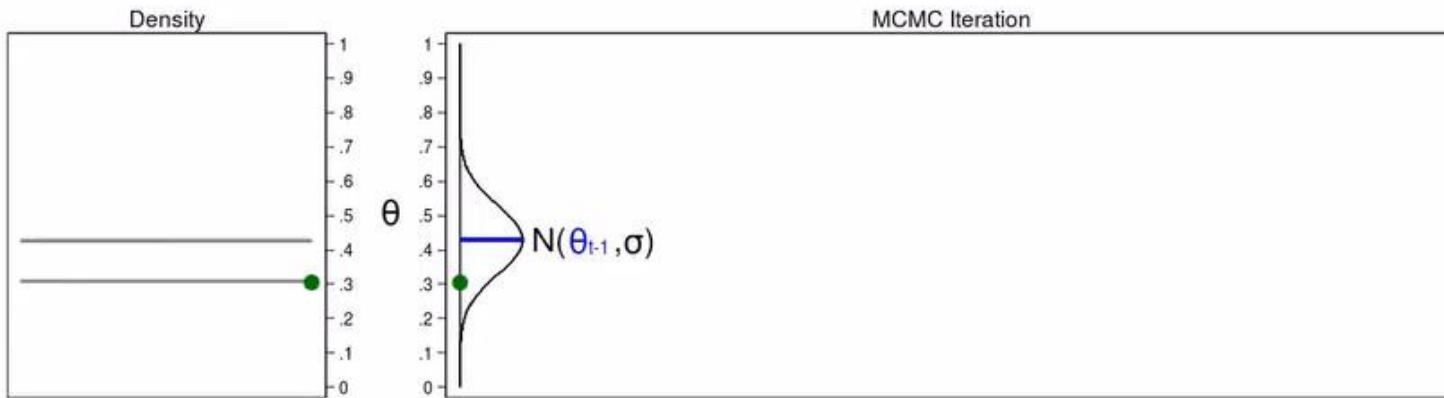
$$\text{Step 1: } r(\theta_{\text{new}}, \theta_{t-1}) = \frac{\text{Posterior}(\theta_{\text{new}})}{\text{Posterior}(\theta_{t-1})} = \frac{\text{Beta}(1,1,0.088) \times \text{Binomial}(10,4, 0.088)}{\text{Beta}(1,1,0.286) \times \text{Binomial}(10,4, 0.286)} = 0.039$$

$$\text{Step 2: Acceptance probability } \alpha(\theta_{\text{new}}, \theta_{t-1}) = \min\{r(\theta_{\text{new}}, \theta_{t-1}), 1\} = \min\{0.039, 1\} = 0.039$$

$$\text{Step 3: Draw } u \sim \text{Uniform}(0,1) = 0.247$$

$$\text{Step 4: If } u < \alpha(\theta_{\text{new}}, \theta_{t-1}) \rightarrow \text{If } 0.247 < 0.039 \quad \begin{array}{l} \text{Then } \theta_t = \theta_{\text{new}} = 0.088 \\ \text{Otherwise } \theta_t = \theta_{t-1} = 0.286 \end{array}$$

MCMC with Metropolis-Hastings



$$\text{Step 1: } r(\theta_{\text{new}}, \theta_{t-1}) = \frac{\text{Posterior}(\theta_{\text{new}})}{\text{Posterior}(\theta_{t-1})} = \frac{\text{Beta}(1,1, 0.306) \times \text{Binomial}(10,4, 0.306)}{\text{Beta}(1,1, 0.429) \times \text{Binomial}(10,4, 0.429)} = 0.834$$

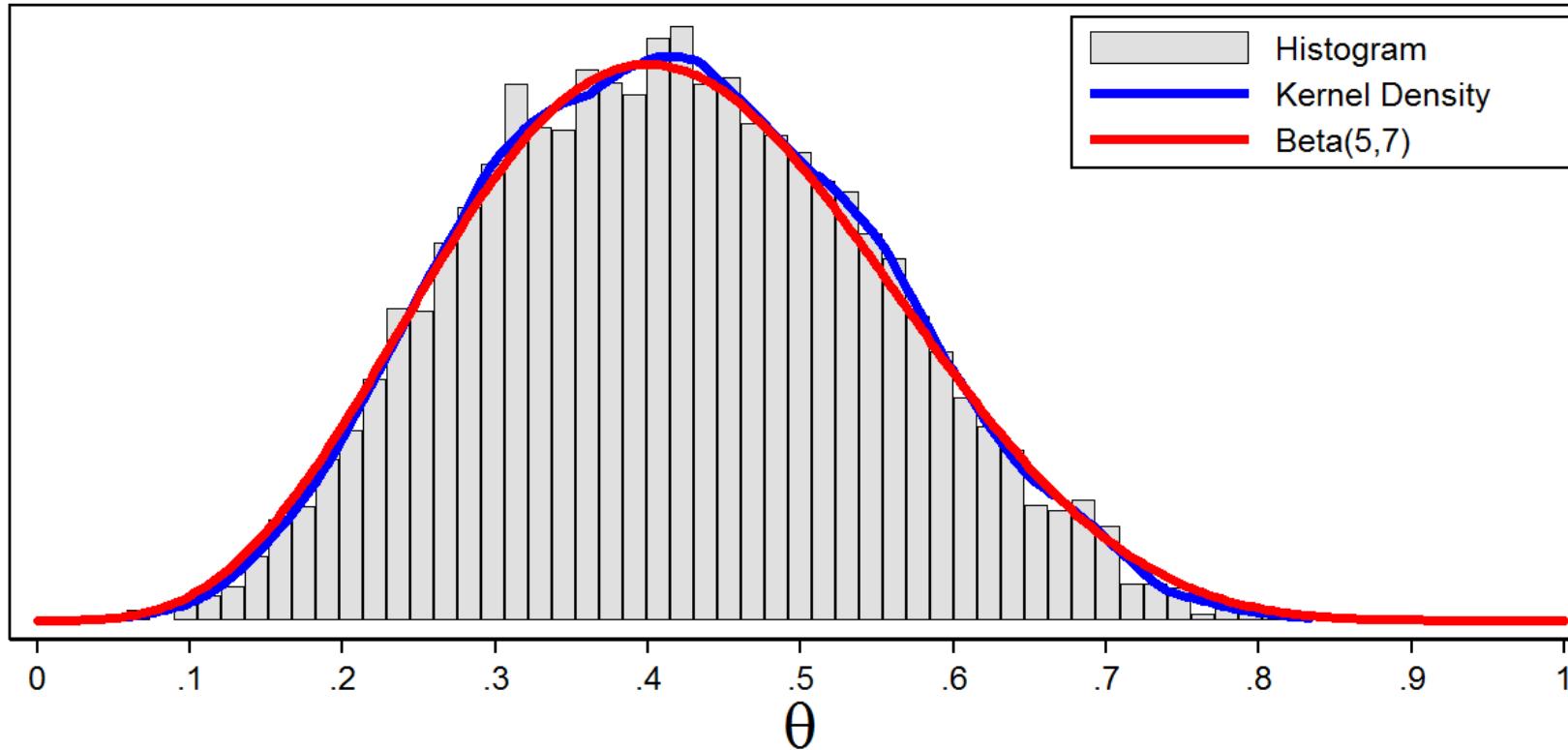
$$\text{Step 2: Acceptance probability } \alpha(\theta_{\text{new}}, \theta_{t-1}) = \min\{r(\theta_{\text{new}}, \theta_{t-1}), 1\} = \min\{0.834, 1\} = 0.834$$

Step 3: Draw $u \sim \text{Uniform}(0,1) = 0.617$

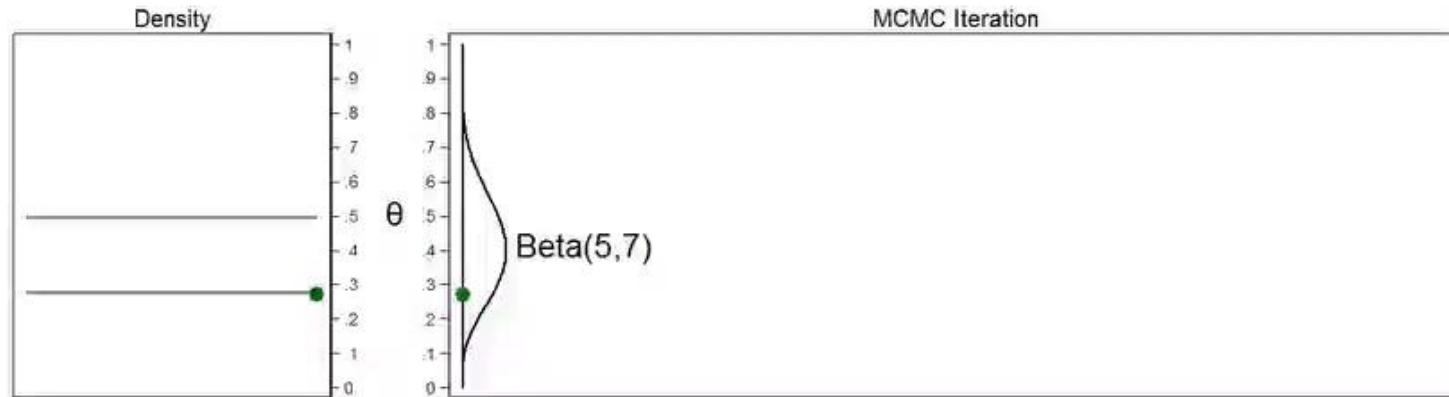
Step 4: If $u < \alpha(\theta_{\text{new}}, \theta_{t-1}) \rightarrow$ If $0.617 < 0.834$ Then $\theta_t = \theta_{\text{new}} = 0.306$
Otherwise $\theta_t = \theta_{t-1} = 0.429$

MCMC with Metropolis-Hastings

Comparison of the MCMC sample and
the theoretical posterior density



MCMC with Gibbs Sampling



Draw $\theta_t \sim \text{Beta}(1,1) \times \text{Binomial}(4,10)$

$\text{Beta}(5,7) = 0.271$

The **bayesmh** command

```
bayesmh heads,                                ///
likelihood(dbernoulli({theta}))    ///
prior({theta}, beta(1,1))
```

```
. bayesmh heads, likelihood(dbernoulli({theta})) prior({theta}, beta(1,1))
```

Burn-in ...

Simulation ...

Model summary

Likelihood:

heads ~ bernoulli({theta})

Prior:

{theta} ~ beta(1,1)

Bayesian Bernoulli model

Random-walk Metropolis-Hastings sampling

MCMC iterations = 12,500

Burn-in = 2,500

MCMC sample size = 10,000

Number of obs = 10

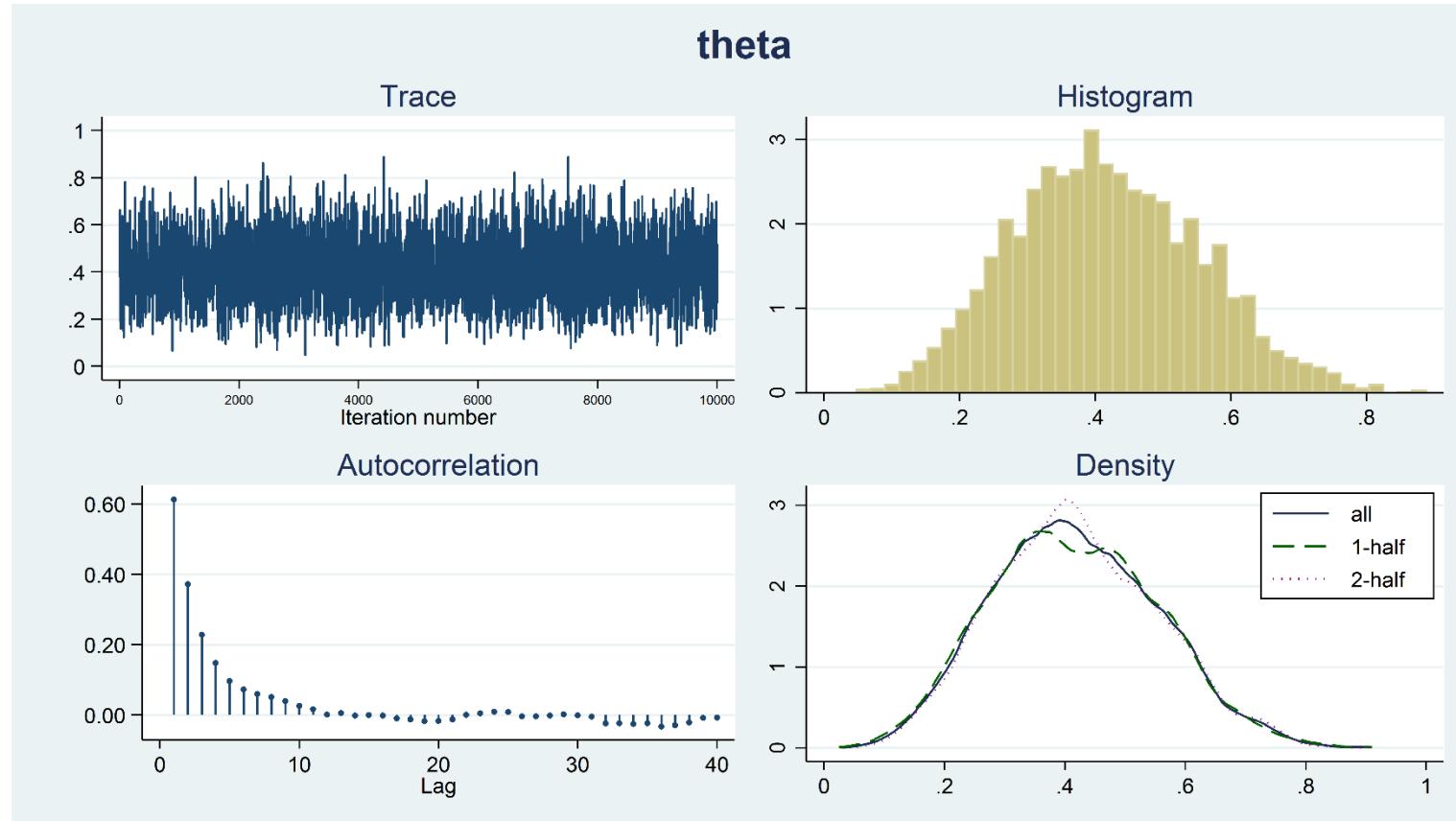
Acceptance rate = .4823

Efficiency = .2291

Log marginal likelihood = -7.8194591

	Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
theta	.4187117	.1342192	.002804	.4152274	.1746616	.6876875

Diagnostic Plots



`bayesgraph diagnostics {theta}`

Outline

- Introduction to Bayesian Analysis
- Coin Toss Example
- Priors, Likelihoods, and Posteriors
- Markov Chain Monte Carlo (MCMC)
- **Bayesian Linear Regression**
- Advantages and Disadvantages of Bayes

Bayesian Linear Regression

```
. webuse nhanes2, clear  
. generate id = _n  
. rename bpsystol sbp  
. describe id sbp highbp healthstatus age sex race
```

variable	storage type	display format	value label	variable label
id	float	%9.0g		Identification Number
sbp	int	%9.0g		Systolic Blood Pressure (mm/Hg)
highbp	byte	%8.0g	highbp	sbp >= 140 or dbp >= 90
healthstatus	byte	%9.0g	hs	Health Status
age	byte	%9.0g		Age (years)
sex	byte	%9.0g	sex	Sex
race	byte	%9.0g	race	Race

We will ignore the sample weights to keep things simple.

$$sbp_i = \beta_0 + \beta_1 age_i + \beta_2 sex_i + e_i$$

. regress sbp age sex

Source	SS	df	MS	Number of obs	=	10,351
Model	1345777.75	2	672888.876	F(2, 10348)	=	1623.51
Residual	4288892.27	10,348	414.465817	Prob > F	=	0.0000
Total	5634670.03	10,350	544.412563	R-squared	=	0.2388
				Adj R-squared	=	0.2387
				Root MSE	=	20.358

sbp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.6530798	.0116249	56.18	0.000	.6302929	.6758668
sex	-4.013619	.4007284	-10.02	0.000	-4.799124	-3.228114
_cons	105.9298	.8447992	125.39	0.000	104.2738	107.5857



. bayes: regress sbp age sex

Burn-in ...

Simulation ...

Model summary

Likelihood:

sbp ~ regress(xb_sbp, {sigma2})

Priors:

{sbp:age sex _cons} ~ normal(0,10000) (1)
{sigma2} ~ igamma(.01,.01)

(1) Parameters are elements of the linear form xb_sbp.

Bayesian linear regression

Random-walk Metropolis-Hastings sampling

MCMC iterations = 12,500
Burn-in = 2,500
MCMC sample size = 10,000
Number of obs = 10,351
Acceptance rate = .3484
Efficiency: min = .07126
avg = .1153
max = .2286

Log marginal likelihood = -45908.106

	Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
sbp						
age	.6529088	.0113266	.000398	.6530755	.630286	.67464
sex	-3.998945	.3923714	.013853	-3.991372	-4.774898	-3.224494
_cons	105.9215	.8241964	.030876	105.917	104.343	107.542
sigma2	414.5504	5.84057	.122166	414.3584	403.5949	426.2577

Note: Default priors are used for model parameters.



bayes: regress - Bayesian linear regression

Model if/in Weights Priors Simulation Blocking Initialization Adaptation Reporting Advanced

Dependent variable: sbp Independent variables: age sex

Suppress constant term

OK Cancel Submit



bayes: regress - Bayesian linear regression

Model if/in Weights Priors Simulation Blocking Initialization Adaptation Reporting Advanced

Use Gibbs sampling to update model parameters

Default priors

Normal prior for coefficients
100 Standard deviation

Inverse-gamma prior for variance components
0.01 Shape
0.01 Scale

Custom priors for model parameters

Press "Create" to define a prior distribution

Show model summary without estimation



bayes: regress - Bayesian linear regression

Model if/in Weights Priors Simulation Blocking Initialization Adaptation Reporting Advanced

10000 MCMC sample size
2500 Number of iterations for the burn-in period of MCMC
1 Thinning interval
Random-number seed

Model parameters to be excluded from simulation results:

OK Cancel Submit



bayes: regress - Bayesian linear regression

Model if/in Weights Priors Simulation Blocking Initialization Adaptation Reporting Advanced

Adaptive MCMC procedure

100 Adaptation interval

25 Maximum number of adaptive iterations

5 Minimum number of adaptive iterations

0.75 Parameter controlling acceptance rate, alpha()

0.8 Parameter controlling proposal covariance, beta()

0 Parameter controlling adaptation rate, gamma()

Target acceptance rate for all blocks of model parameters

0.01 Tolerance for acceptance rate

2.38 Initial multiplier for the scale factor for all blocks

Scale matrix for initial proposal covariance

OK Cancel Submit

bayes options

```
. bayes: regress sbp age sex, noheader
```

Burn-in ...

Simulation ...

	Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
sbp						
age	.6530982	.0110625	.000405	.6533669	.6318849	.6737548
sex	-4.010928	.3865944	.013339	-4.015244	-4.734603	-3.241869
_cons	105.9341	.8156657	.030091	105.9519	104.2668	107.5468
sigma2	414.5182	5.636057	.11081	414.3449	403.8492	425.8631

bayes options

. bayes, rseed(15): regress sbp age sex, noheader

Burn-in ...

Simulation ...

	Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
sbp						
age	.6529088	.0113266	.000398	.6530755	.630286	.67464
sex	-3.998945	.3923714	.013853	-3.991372	-4.774898	-3.224494
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